



## The residual based interactive least squares algorithms and simulation studies

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### ABSTRACT

This paper presents a two-stage least squares based iterative algorithm, a residual based interactive least squares algorithm and a residual based recursive least squares algorithm for identifying controlled autoregressive moving average (C-ARMA) models. The simulation studies indicate that the proposed algorithms can effectively estimate the parameters of the C-ARMA models.

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### 1. Introduction

The time series contains three basic models: autoregressive (AR) model, moving average (MA) model and autoregressive moving average (ARMA) model. This paper considers the least squares identification problem of controlled autoregressive moving average (C-ARMA) model.

Compared with least squares (LS) algorithm, the stochastic gradient (SG) algorithm has small computational load but slow convergence rate [1]. Recently, Ding, Yang and Liu analyzed the consistency of the multivariable SG algorithm [2]; Ding and Chen presented a hierarchical SG algorithm for multivariable systems [3] and an auxiliary model based SG algorithm for dual-rate systems [4]; Ding et al. studied the performances of the SG algorithms for dual-rate systems based on the polynomial transformation technique [5,6]. In order to improve the convergence rate of the SG algorithm, Ding and Chen developed a multi-innovation SG identification algorithm for linear regression model [7] and an extended stochastic gradient algorithm with a forgetting factor for Hammerstein nonlinear systems [8]; Ding and Wang discussed the gradient based identification algorithm for Hammerstein–Wiener ARMAX systems [9]. Finally, Zhang, Ding and Shi presented a multi-innovation SG parameter estimation based self-tuning control algorithm [10].

Because of the fast convergence rate of the least squares identification, it has received much attention in many areas, including signal processing [11], system identification and parameter estimation [12–19], adaptive control [5,20,21]. For example, Ding and Chen presented a hierarchical LS algorithm for multivariable systems [22], whose consistency was studied in [23]; Ding and Chen proposed an auxiliary model based LS algorithm for dual-rate systems [24]; Ding, Liu and Shi studied the performances of the polynomial transform based LS algorithm for dual-rate systems [25].

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### 2. The system description

Consider a controlled autoregressive moving average (C-ARMA) model,

$$y(t) = B(z)u(t) + \frac{D(z)}{C(z)}v(t), \tag{1}$$

where  $\{y(t)\}$  is the observation output,  $\{u(t)\}$  is the control input,  $\{v(t)\}$  is a white noise sequence with zero mean  $E[v(t)] = 0$  and variance  $\sigma^2 := E[v^2(t)]$  ( $E$  denotes the expectation operator) and independent of  $u(t)$ ,  $z^{-1}$  represents the unit backward shift operator:  $z^{-1}y(t) = y(t - 1)$ , and  $B(z)$ ,  $C(z)$  and  $D(z)$  are polynomials in  $z^{-1}$  with

$$\begin{aligned} B(z) &= b_1z^{-1} + b_2z^{-2} + \dots + b_{n_b}z^{-n_b}, \\ C(z) &= 1 + c_1z^{-1} + c_2z^{-2} + \dots + c_{n_c}z^{-n_c}, \\ D(z) &= 1 + d_1z^{-1} + d_2z^{-2} + \dots + d_{n_d}z^{-n_d}. \end{aligned}$$

Assume that the orders  $n_b$ ,  $n_c$  and  $n_d$  are known. Let

$$w(t) := \frac{D(z)}{C(z)}v(t). \tag{2}$$

$w(t)$  is a correlated noise with zero mean value and variance:

$$\sigma_w^2 = \frac{\sigma^2}{2\pi j} \oint_{\gamma} \frac{D(z) D(z^{-1})}{C(z) C(z^{-1})} \frac{dz}{z}, \quad j = \sqrt{-1},$$

where  $\gamma$  is a unit circle with  $|\gamma| = 1$ .

Define the parameter vectors  $\theta$  and  $\vartheta$  and information vectors  $\phi(t)$ ,  $\phi(t)$  and  $\psi(t)$  as

$$\begin{aligned} \theta &:= \begin{bmatrix} \mathbf{b} \\ \vartheta \end{bmatrix} \in \mathbb{R}^{n_b+n_c+n_d}, & \phi(t) &:= \begin{bmatrix} \phi(t) \\ \psi(t) \end{bmatrix} \in \mathbb{R}^{n_b+n_c+n_d}, \\ \mathbf{b} &:= [b_1, b_2, \dots, b_{n_b}]^T \in \mathbb{R}^{n_b}, \\ \vartheta &:= [c_1, c_2, \dots, c_{n_c}, d_1, d_2, \dots, d_{n_d}]^T \in \mathbb{R}^{n_c+n_d}, \\ \phi(t) &:= [u(t-1), u(t-2), \dots, u(t-n_b)]^T \in \mathbb{R}^{n_b}, \\ \psi(t) &:= [-w(t-1), -w(t-2), \dots, -w(t-n_c), v(t-1), v(t-2), \dots, v(t-n_d)]^T \in \mathbb{R}^{n_c+n_d}, \end{aligned} \tag{3}$$

where the superscript T represents the transpose of a matrix or vector.

From (2) and (1), we have the following identification models:

$$w(t) = \psi^T(t)\vartheta + v(t), \tag{4}$$

$$\begin{aligned} y(t) &= \phi^T(t)\mathbf{b} + w(t) \\ &= \phi^T(t)\mathbf{b} + \psi^T(t)\vartheta + v(t) \end{aligned} \tag{5}$$

$$= \varphi^T(t)\theta + v(t). \tag{6}$$

Note that  $y(t)$  and  $\phi(t)$  are measured but  $w(t)$ ,  $\psi(t)$  and  $v(t)$  are unknown (unmeasured). The objective of this paper is to present the residual based identification algorithms to estimate the parameter vectors  $\theta$  or  $\mathbf{b}$  and  $\vartheta$  from the observation data  $\{u(t), y(t)\}$  or  $\{y(t), \phi(t)\}$ .

### 3. The least squares estimate of $\mathbf{b}$

Define the stacked vectors and matrices:

$$\begin{aligned} \mathbf{Y}(t) &:= \begin{bmatrix} y(1) \\ y(2) \\ \vdots \\ y(t) \end{bmatrix} \in \mathbb{R}^t, & \mathbf{H}(t) &:= \begin{bmatrix} \phi^T(1) \\ \phi^T(2) \\ \vdots \\ \phi^T(t) \end{bmatrix} \in \mathbb{R}^{t \times n_b}, \\ \mathbf{W}(t) &:= \begin{bmatrix} w(1) \\ w(2) \\ \vdots \\ w(t) \end{bmatrix} \in \mathbb{R}^t, & \Psi(t) &:= \begin{bmatrix} \psi^T(1) \\ \psi^T(2) \\ \vdots \\ \psi^T(t) \end{bmatrix} \in \mathbb{R}^{t \times (n_c+n_d)}, & \mathbf{V}(t) &:= \begin{bmatrix} v(1) \\ v(2) \\ \vdots \\ v(t) \end{bmatrix} \in \mathbb{R}^t. \end{aligned} \tag{7}$$

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