



The residual based interactive least squares algorithms and simulation studies

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ARTICLE INFO

Article history:

Received 6 March 2008

Received in revised form 15 February 2009

Accepted 27 February 2009

Keywords:

Recursive identification

Parameter estimation

Least squares

Controlled autoregressive moving average (C-ARMA) models

AR models

MA models

ABSTRACT

This paper presents a two-stage least squares based iterative algorithm, a residual based interactive least squares algorithm and a residual based recursive least squares algorithm for identifying controlled autoregressive moving average (C-ARMA) models. The simulation studies indicate that the proposed algorithms can effectively estimate the parameters of the C-ARMA models.

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1. Introduction

The time series contains three basic models: autoregressive (AR) model, moving average (MA) model and autoregressive moving average (ARMA) model. This paper considers the least squares identification problem of controlled autoregressive moving average (C-ARMA) model.

Compared with least squares (LS) algorithm, the stochastic gradient (SG) algorithm has small computational load but slow convergence rate [1]. Recently, Ding, Yang and Liu analyzed the consistency of the multivariable SG algorithm [2]; Ding and Chen presented a hierarchical SG algorithm for multivariable systems [3] and an auxiliary model based SG algorithm for dual-rate systems [4]; Ding et al. studied the performances of the SG algorithms for dual-rate systems based on the polynomial transformation technique [5,6]. In order to improve the convergence rate of the SG algorithm, Ding and Chen developed a multi-innovation SG identification algorithm for linear regression model [7] and an extended stochastic gradient algorithm with a forgetting factor for Hammerstein nonlinear systems [8]; Ding and Wang discussed the gradient based identification algorithm for Hammerstein–Wiener ARMAX systems [9]. Finally, Zhang, Ding and Shi presented a multi-innovation SG parameter estimation based self-tuning control algorithm [10].

Because of the fast convergence rate of the least squares identification, it has received much attention in many areas, including signal processing [11], system identification and parameter estimation [12–19], adaptive control [5,20,21]. For example, Ding and Chen presented a hierarchical LS algorithm for multivariable systems [22], whose consistency was studied in [23]; Ding and Chen proposed an auxiliary model based LS algorithm for dual-rate systems [24]; Ding, Liu and Shi studied the performances of the polynomial transform based LS algorithm for dual-rate systems [25].

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2. The system description

Consider a controlled autoregressive moving average (C-ARMA) model,

$$y(t) = B(z)u(t) + \frac{D(z)}{C(z)}v(t), \tag{1}$$

where $\{y(t)\}$ is the observation output, $\{u(t)\}$ is the control input, $\{v(t)\}$ is a white noise sequence with zero mean $E[v(t)] = 0$ and variance $\sigma^2 := E[v^2(t)]$ (E denotes the expectation operator) and independent of $u(t)$, z^{-1} represents the unit backward shift operator: $z^{-1}y(t) = y(t - 1)$, and $B(z)$, $C(z)$ and $D(z)$ are polynomials in z^{-1} with

$$\begin{aligned} B(z) &= b_1z^{-1} + b_2z^{-2} + \dots + b_{n_b}z^{-n_b}, \\ C(z) &= 1 + c_1z^{-1} + c_2z^{-2} + \dots + c_{n_c}z^{-n_c}, \\ D(z) &= 1 + d_1z^{-1} + d_2z^{-2} + \dots + d_{n_d}z^{-n_d}. \end{aligned}$$

Assume that the orders n_b , n_c and n_d are known. Let

$$w(t) := \frac{D(z)}{C(z)}v(t). \tag{2}$$

$w(t)$ is a correlated noise with zero mean value and variance:

$$\sigma_w^2 = \frac{\sigma^2}{2\pi j} \oint_{\gamma} \frac{D(z) D(z^{-1})}{C(z) C(z^{-1})} \frac{dz}{z}, \quad j = \sqrt{-1},$$

where γ is a unit circle with $|\gamma| = 1$.

Define the parameter vectors θ and ϑ and information vectors $\phi(t)$, $\phi(t)$ and $\psi(t)$ as

$$\begin{aligned} \theta &:= \begin{bmatrix} \mathbf{b} \\ \vartheta \end{bmatrix} \in \mathbb{R}^{n_b+n_c+n_d}, & \phi(t) &:= \begin{bmatrix} \phi(t) \\ \psi(t) \end{bmatrix} \in \mathbb{R}^{n_b+n_c+n_d}, \\ \mathbf{b} &:= [b_1, b_2, \dots, b_{n_b}]^T \in \mathbb{R}^{n_b}, \\ \vartheta &:= [c_1, c_2, \dots, c_{n_c}, d_1, d_2, \dots, d_{n_d}]^T \in \mathbb{R}^{n_c+n_d}, \\ \phi(t) &:= [u(t-1), u(t-2), \dots, u(t-n_b)]^T \in \mathbb{R}^{n_b}, \\ \psi(t) &:= [-w(t-1), -w(t-2), \dots, -w(t-n_c), v(t-1), v(t-2), \dots, v(t-n_d)]^T \in \mathbb{R}^{n_c+n_d}, \end{aligned} \tag{3}$$

where the superscript T represents the transpose of a matrix or vector.

From (2) and (1), we have the following identification models:

$$w(t) = \psi^T(t)\vartheta + v(t), \tag{4}$$

$$\begin{aligned} y(t) &= \phi^T(t)\mathbf{b} + w(t) \\ &= \phi^T(t)\mathbf{b} + \psi^T(t)\vartheta + v(t) \end{aligned} \tag{5}$$

$$= \varphi^T(t)\theta + v(t). \tag{6}$$

Note that $y(t)$ and $\phi(t)$ are measured but $w(t)$, $\psi(t)$ and $v(t)$ are unknown (unmeasured). The objective of this paper is to present the residual based identification algorithms to estimate the parameter vectors θ or \mathbf{b} and ϑ from the observation data $\{u(t), y(t)\}$ or $\{y(t), \phi(t)\}$.

3. The least squares estimate of \mathbf{b}

Define the stacked vectors and matrices:

$$\begin{aligned} \mathbf{Y}(t) &:= \begin{bmatrix} y(1) \\ y(2) \\ \vdots \\ y(t) \end{bmatrix} \in \mathbb{R}^t, & \mathbf{H}(t) &:= \begin{bmatrix} \phi^T(1) \\ \phi^T(2) \\ \vdots \\ \phi^T(t) \end{bmatrix} \in \mathbb{R}^{t \times n_b}, \\ \mathbf{W}(t) &:= \begin{bmatrix} w(1) \\ w(2) \\ \vdots \\ w(t) \end{bmatrix} \in \mathbb{R}^t, & \Psi(t) &:= \begin{bmatrix} \psi^T(1) \\ \psi^T(2) \\ \vdots \\ \psi^T(t) \end{bmatrix} \in \mathbb{R}^{t \times (n_c+n_d)}, & \mathbf{V}(t) &:= \begin{bmatrix} v(1) \\ v(2) \\ \vdots \\ v(t) \end{bmatrix} \in \mathbb{R}^t. \end{aligned} \tag{7}$$

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