The Situational Triggers of Aggressive Responses scale in five countries: Factor structure and country clustering solutions

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Abstract

Much research has focused on the detection of factors that might lead to aggression based on evolutionary arguments, trait approaches, psychosocial models and cognition/information processing. There may be trait differences in the way that individuals respond to environmental triggers, thus, aggression may occur as the expression of a trait during specific situations. Such specific situations have been described (Lawrence, 2006) as "situational triggers" and are assessed through the STAR (Situational Triggers of Aggressive Responses) scale which is comprised of two main situational prompts, Provocations and Frustrations. In Study 1 (N = 328 Greek university students), confirmatory factor analysis modeling confirmed the STAR scale structure with minor fluctuations. In Study 2, using data from the UK, Poland, Korea and USA, as well as an additional sample of Greek participants (N = 1219), we tested the STAR for factor equivalence levels, aiming at an overall factor structure. The scale structure was confirmed across countries with levels of factor equivalence being satisfactory, although some within-factor collinearity was observed. A clusters-of-countries approach was thus implemented for further testing within each cluster. Overall, the stability and validity levels of the STAR structure and its cross-cultural application were verified with possible considerations of country-sets being the units of future analysis.

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1. Introduction

Estimating when a person is likely to behave aggressively is a key step in intervening to prevent aggression (Barling, Dupré, & Kelloway, 2009). Research on this topic has focused on several different levels of analysis. First, evolutionary arguments have sought to explain aggressive behavior as a function of Darwin’s sexual selection (e.g., Archer, 2009), which may be subject to geographical and culture influences. Second, trait approaches have examined which individuals are more likely to act aggressively (Buss & Perry, 1992), narcissism (Bushman et al., 2005), impulsivity (Cross, Copping, & Campbell, 2011) and hostile attribution (Crick & Dodge, 1996) all increasing the likelihood that a person will behave aggressively (Bettencourt, Talley, Benjamin, & Valentine, 2006; Lawrence & Hodgkins, 2009; Ortiz & Raine, 2004; Wilkowski, Robinson, & Troop-Gordon, 2010). Finally, psychosocial models have pointed to environmental influences on aggression. These environmental influences can be broad, such as the relative differences in approval for aggressive behavior across cultures or by sex (Thanzami & Archer, 2005), or they can be more specific, such as the presence of provocations (Anderson & Bushman, 2002), frustrations (Berkowitz, 2008), and additional demands on self-control (Stucke & Baumeister, 2006). Thus, much of the research has focused on the detection of factors that might lead to aggression, with both personality factors such as aggressive/angry traits, and situational factors such as frustration, stress, or state anger (McCurdy, 2005; Sprague, Verona, Kalkhoff, & Kilmer, 2008), playing a role. While this research suggests that both dispositional and situational triggers may increase the likelihood of aggressive responding, more subtle individual differences in cognition or information processing (e.g., schemata; Milner et al., 2011) also play a key role in influencing the propensity for aggressive behavior. For example, individual differences in the tendency to attribute hostile intent are associated with variations in aggressive responding (see Crick & Dodge, 1994 for a review). These data are consistent with more recent research, which demonstrates significant individual differences in reactivity to hostile cues (e.g., Lawrence & Hodgkins, 2009; Robinson & Wilkowski, 2010).

In fact, it has been proposed that there may be trait differences in the way that individuals respond to these environmental triggers (Lawrence, 2006). Thus, the tendency for aggression in response to specific situations has a trait-like quality, with individuals varying in their inclination to act aggressively in response to situational...
triggers. Specific situations that often trigger aggression have been described as “situational triggers” (Lawrence, 2006) and have served as the basis for an instrument designed to “predict individual differences in the kinds of events and antecedents that make people feel aggressive” (p. 242). Based on these theoretical grounds, the Situational Triggers of Aggressive Responses (STAR) scale was created in order to assess aggression under two main situational prompts, Provocations and Frustrations. Through a series of studies, Lawrence (2006) devised and tested a set of items, scored by youth and adults (16 years or older) reflecting individuals’ self-reported propensity to respond aggressively to various triggers, as related to Sensitivity to Frustrations (SF) and Sensitivity to Provocations (SP) (further details on the STAR scale are given in Section 2.2). The extent to which the STAR scale’s structure is reproducible across different countries and cultures has been examined to a limited extent only. To date, apart from the initial UK studies, the STAR has also been used in Germany (Bondü & Richter, 2016) and at least two other countries, Poland and Greece, with research supporting convergent validity for the STAR in Leung, 1997). The proximity and continent. This way, we may better understand the structure across countries that are diverse in terms of geographic proximity and continent. This way, we may better understand the cross-cultural similarities and differences in aggression mechanisms, taking other situation-specific correlates into consideration as well.

With these in mind, the main purpose of the current two studies is to test the factor structure of the STAR scale. For a scale to be readily available for use in different cultures, it must support various levels of equivalence starting with factor structure equivalence (Van de Vijver & Leung, 1997). The first study is a confirmatory factor analysis approach applied to a Greek sample of 328 university students. This first study served as a pilot one and a guide for the second study where data (N = 1219) from the UK, Poland, South Korea and the USA, as well as an additional sample of Greek participants were examined. The aims of the second study were to test the STAR scale within these five countries for factor equivalence levels, and if possible, derive an overall factor structure for these countries.

2. Study 1

2.1. Purpose and sample characteristics

The purpose of this study was to confirm the existence of the two STAR dimensions (Sensitivity to Provocations and Sensitivity to Frustrations) in a Greek sample of university students. The sample consisted of 328 students, 203 of which studied psychology, 106 of which studied sciences, and 15 of which studied medicine. Approximately 70% of these students (N = 229) were females and the mean age for the total sample was 21.34 years (SD = 3.34). These data were collected during the 2013–2014 academic year.

2.2. Materials and procedure

The STAR scale consists of 22 items, 10 of which comprise the Sensitivity to Frustrations (SF) factor and 12 of which comprise the Sensitivity to Provocations (SP) factor; all items are scored on a 5-point Likert-type scale and high-scores reflect individuals’ self-reported propensity to respond aggressively to the two triggers. Scale construction and preliminary psychometric data for this instrument were based on UK samples. Both scales demonstrate good reliability (Cronbach’s α = .82 and .80 for SF and SP, respectively) and convergent validity (Lawrence, 2006; Lawrence & Hodgkins, 2009; Lawrence & Hutchinson, 2013a; Lawrence & Hutchinson, 2013b).

The STAR scale was first translated into Greek and was then back-translated into English by two experts. A few language modifications were necessary and cultural specificities were considered (so as to avoid sex-discriminating language and to convey the real cultural meaning of items as much as possible). The instrument was administered (having attained informed consent) to several small groups of university students who were awaiting a lecture.

2.3. Results

We first conducted confirmatory factor analysis (Jöreskog & Sörbom, 1996; Schreiber, Nora, Stage, Barlow, & King, 2006) on the 22 STAR items with respect to the original theory structure (Lawrence, 2006). The following criteria and indices were considered: Normal theory weighted least squares chi-square ($\chi^2$) and $\chi^2$/df. Comparative Fit Index (CFI), Goodness of Fit index (GFI), Root Mean Square Error of Approximation (RMSEA), and for the comparison across models, Tucker-Lewis Index (TLI) and $\Delta_2^2$ along with its significance levels. In an initial two-factor solution, goodness of fit indices did not reach acceptable levels.

A closer look revealed some distribution irregularities that were possibly affecting correlation (Pearson’s r) magnitudes. Recognizing that a non-parametric approach might improve the potential and precision of the correlation indices, a Fisher’s-z transformation (Mylonas, 2009; Mylonas, Veligekas, Gari, & Kontaxopoulou, 2012; Steiger, 1980; Zajenkowska et al., 2014) was used to compare the Pearson’s r indices with Spearman’s Rho and Kendall’s Tau-b indices. Kendall’s Tau-b was used in all subsequent analyses because Kendall’s Tau-b coefficients did not statistically differ from Pearson’s r indices; for the overall sample only 10 out of the 231 pairs of coefficients were different at the .05 significance level (approx. 4%). Thus, the analysis on Kendall Tau-b is fully justified in this research as these coefficients only add/unmask information and do not alter the overall intercorrelation matrix. Kendall’s approach is preferable to the Spearman’s Rho to address statistical ties in the data (Howell, 1987). Nunnally and Bernstein (1994), in their discussion of measures of linear relation, suggest that “None has achieved the prominence of r because none fits as neatly into the mathematics of a general psychometric theory […] The closest to an exception is Kendall’s tau...” (pp. 124–125). Apart from this assertion that Kendall’s Tau coefficient can fit best in a general psychometric theory and this makes it a very good Pearson’s r substitute, Nunnally and Bernstein explain that psychometricians should be cautious with the use of Pearson’s r when equal intervals are not given, despite any continuity that may seem to exist in the data (i.e., Likert-type measurement scales).

A series of CFA models were applied to the data (Table 1). The independence model was easily rejected with a $\chi^2$/df index reaching as high as 6.00. A unifactorial model was much better, but RMSEA remained at high levels and CFI was very low. The two-factor model showed improved goodness of fit, but RMSEA still remained rather high, and GFI and CFI did not improve much; however, there was some model improvement with respect to the unifactorial solution ($\Delta_2^2 = .20$).

A modified two-factor model reached acceptable levels with RMSEA being even lower than .05, GFI exceeding .90, CFI approaching .90, and $\chi^2$/df dropping below 2 (despite the notorious and unavoidable $\chi^2$ significance). Model improvement with respect to the unifactorial model was .51, a satisfactory TLI outcome. This final model in Table 1 is a modified two-factor model for which seven error covariances—strictly within-in factors—were estimated, allowing for some collinearity levels within factors (factors remaining independent). While this model was acceptable, it is noteworthy that it was the Sensitivity to Provocations factor which suffered the most from collinearity problems.
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